

## **GCE**

# **Mathematics (MEI)**

Unit 4769: Statistics 4

**Advanced GCE** 

Mark Scheme for June 2015

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This mark scheme is published as an aid to teachers and students, to indicate the requirements of the examination. It shows the basis on which marks were awarded by examiners. It does not indicate the details of the discussions which took place at an examiners' meeting before marking commenced.

All examiners are instructed that alternative correct answers and unexpected approaches in candidates' scripts must be given marks that fairly reflect the relevant knowledge and skills demonstrated.

Mark schemes should be read in conjunction with the published question papers and the report on the examination.

OCR will not enter into any discussion or correspondence in connection with this mark scheme.

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### **Annotations and abbreviations**

Annotation in scoris	Meaning
√and <b>x</b>	
BOD	Benefit of doubt
FT	Follow through
ISW	Ignore subsequent working
M0, M1	Method mark awarded 0, 1
A0, A1	Accuracy mark awarded 0, 1
B0, B1	Independent mark awarded 0, 1
SC	Special case
۸	Omission sign
MR	Misread
Highlighting	
Other abbreviations in	Meaning
mark scheme	
E1	Mark for explaining
U1	Mark for correct units
G1	Mark for a correct feature on a graph
M1 dep*	Method mark dependent on a previous mark, indicated by *
cao	Correct answer only
oe	Or equivalent
rot	Rounded or truncated
soi	Seen or implied
www	Without wrong working

#### Subject-specific Marking Instructions for GCE Mathematics (MEI) Statistics strand

a Annotations should be used whenever appropriate during your marking.

The A, M and B annotations must be used on your standardisation scripts for responses that are not awarded either 0 or full marks. It is vital that you annotate standardisation scripts fully to show how the marks have been awarded.

For subsequent marking you must make it clear how you have arrived at the mark you have awarded.

An element of professional judgement is required in the marking of any written paper. Remember that the mark scheme is designed to assist in marking incorrect solutions. Correct solutions leading to correct answers are awarded full marks but work must not be judged on the answer alone, and answers that are given in the question, especially, must be validly obtained; key steps in the working must always be looked at and anything unfamiliar must be investigated thoroughly.

Correct but unfamiliar or unexpected methods are often signalled by a correct result following an *apparently* incorrect method. Such work must be carefully assessed. When a candidate adopts a method which does not correspond to the mark scheme, award marks according to the spirit of the basic scheme; if you are in any doubt whatsoever (especially if several marks or candidates are involved) you should contact your Team Leader.

c The following types of marks are available.

#### M

A suitable method has been selected and *applied* in a manner which shows that the method is essentially understood. Method marks are not usually lost for numerical errors, algebraic slips or errors in units. However, it is not usually sufficient for a candidate just to indicate an intention of using some method or just to quote a formula; the formula or idea must be applied to the specific problem in hand, eg by substituting the relevant quantities into the formula. In some cases the nature of the errors allowed for the award of an M mark may be specified.

#### Α

Accuracy mark, awarded for a correct answer or intermediate step correctly obtained. Accuracy marks cannot be given unless the associated Method mark is earned (or implied). Therefore M0 A1 cannot ever be awarded.

#### В

Mark for a correct result or statement independent of Method marks.

#### F

A given result is to be established or a result has to be explained. This usually requires more working or explanation than the establishment of an unknown result.

Unless otherwise indicated, marks once gained cannot subsequently be lost, eg wrong working following a correct form of answer is ignored. Sometimes this is reinforced in the mark scheme by the abbreviation isw. However, this would not apply to a case where a candidate passes through the correct answer as part of a wrong argument.

- When a part of a question has two or more 'method' steps, the M marks are in principle independent unless the scheme specifically says otherwise; and similarly where there are several B marks allocated. (The notation 'dep \*' is used to indicate that a particular mark is dependent on an earlier, asterisked, mark in the scheme.) Of course, in practice it may happen that when a candidate has once gone wrong in a part of a question, the work from there on is worthless so that no more marks can sensibly be given. On the other hand, when two or more steps are successfully run together by the candidate, the earlier marks are implied and full credit must be given.
- The abbreviation ft implies that the A or B mark indicated is allowed for work correctly following on from previously incorrect results. Otherwise, A and B marks are given for correct work only differences in notation are of course permitted. A (accuracy) marks are not given for answers obtained from incorrect working. When A or B marks are awarded for work at an intermediate stage of a solution, there may be various alternatives that are equally acceptable. In such cases, exactly what is acceptable will be detailed in the mark scheme rationale. If this is not the case please consult your Team Leader.

Sometimes the answer to one part of a question is used in a later part of the same question. In this case, A marks will often be 'follow through'. In such cases you must ensure that you refer back to the answer of the previous part question even if this is not shown within the image zone. You may find it easier to mark follow through questions candidate-by-candidate rather than question-by-question.

- Wrong or missing units in an answer should not lead to the loss of a mark unless the scheme specifically indicates otherwise. Candidates are expected to give numerical answers to an appropriate degree of accuracy, with 3 significant figures often being the norm. Small variations in the degree of accuracy to which an answer is given (e.g. 2 or 4 significant figures where 3 is expected) should not normally be penalised, while answers which are grossly over- or under-specified should normally result in the loss of a mark. The situation regarding any particular cases where the accuracy of the answer may be a marking issue should be detailed in the mark scheme rationale. If in doubt, contact your Team Leader.
- g Rules for replaced work

If a candidate attempts a question more than once, and indicates which attempt he/she wishes to be marked, then examiners should do as the candidate requests.

If there are two or more attempts at a question which have not been crossed out, examiners should mark what appears to be the last (complete) attempt and ignore the others.

- NB Follow these maths-specific instructions rather than those in the assessor handbook.
- h For a *genuine* misreading (of numbers or symbols) which is such that the object and the difficulty of the question remain unaltered, mark according to the scheme but following through from the candidate's data. A penalty is then applied; 1 mark is generally appropriate, though this may differ for some units. This is achieved by withholding one A mark in the question.

Note that a miscopy of the candidate's own working is not a misread but an accuracy error.

Que	estion	Answer	Marks	Guidance
1 (	(i)	The integral is $-\exp(-x^2/a)$ , limits 0 and infinity. (Which evaluates to $0 - (-1) = 1$ .)	B1	Answer given
(1	(ii)	$E(X^{2}) = \int_{0}^{\infty} \frac{2}{a} x^{3} \exp\left(-\frac{x^{2}}{a}\right) dx  or  \int_{0}^{\infty} x^{2} \frac{2}{a} x \exp\left(-\frac{x^{2}}{a}\right) dx$	[1] M1	
		$= \left[ x^2 \left( -\exp\left(-\frac{x^2}{a}\right) \right) \right]_0^{\infty} + \int_0^{\infty} 2x \exp\left(-\frac{x^2}{a}\right) dx$	M1A1	M1 parts
		=0+a=a	A1	Cao
		$E(X^4) = \int_0^\infty \frac{2}{a} x^5 \exp\left(-\frac{x^2}{a}\right) dx \text{ or } \int_0^\infty x^4 \frac{2}{a} x \exp\left(-\frac{x^2}{a}\right) dx$	M1	
		$= \left[ x^4 \left( -\exp\left(-\frac{x^2}{a}\right) \right) \right]_0^{\infty} + \int_0^{\infty} 4x^3 \exp\left(-\frac{x^2}{a}\right) dx$	A1	Answer given so working
		$= 0 + 2a \times a = 2a^2$	A1	must be convincing
			[7]	
	(iii)	Likelihood is $\left(\frac{2}{a}\right)^n \prod X_i \exp\left(-\frac{1}{a}\sum X_i^2\right)$	M1A1A1	M1 for recognisable attempt to obtain the likelihood. A1 for $\prod X_i$ . A1 for the exp fn.
		(log likelihood as given) Differentiate log likelihood	M1	
		to obtain $-\frac{n}{a} + \frac{1}{a^2} \sum X_i^2$	A1	
		Set this to zero to obtain MLE(a) = $\frac{1}{n}\sum X_i^2$	M1A1	Justification of a maximum
		n	[7]	not required

Question	Answer	Marks	Guidance
(iv)	$E(MLE(a)) = \frac{1}{n} \sum E(X_i^2) = \frac{1}{n} n a = a$	В1	
	That is, the MLE is unbiased	E1	Seen or very clearly implied
	$\operatorname{Var}(\operatorname{MLE}(a)) = \frac{1}{n^2} \sum \operatorname{Var}(X_i^2) = \frac{1}{n^2} \sum \left( \operatorname{E}(X_i^4) - \left( \operatorname{E}(X_i^2)^2 \right) \right)$	M1M1	
	$=\frac{1}{n^2}n(2a^2-a^2)=\frac{a^2}{n}$	A1	
	$-\frac{1}{n^2}n(2\alpha - \alpha ) - \frac{1}{n}$	[5]	
(v)	Maximum likelihood estimate (also unbiased) of a is 1.471	B1	Explanation not required
	with estimated standard error $(a / \sqrt{n})$ 0.1471.	B1	if calculations are correct
	95% CI is $1.471 \pm 1.960 \times 0.1471 = 1.471 \pm 0.288$	M1A1	Accept 2 for 1.960
		[4]	
		[24]	

(	Questi	on Answer	Marks	Guidance
2	(i)	$\mathbf{M}_{Y}(\theta) = \mathbf{E}(\mathbf{e}^{\theta Y}) = \int_{0}^{\infty} \mathbf{e}^{\theta y} \frac{1}{\sqrt{2\pi y}} \mathbf{e}^{-\frac{1}{2}y} dy$	M1, A1	
		$=\int_0^\infty \frac{1}{\sqrt{2\pi y}} e^{-\frac{1}{2}y(1-2\theta)} dy$	A1	
		Substitute $u = y (1 - 2\theta)$ , $du = dy (1 - 2\theta)$ to obtain	M1	
		$\int_0^\infty \frac{\sqrt{1-2\theta}}{\sqrt{2\pi u}} e^{-\frac{1}{2}u} \frac{1}{1-2\theta} du$	A1	
		$= (1 - 2\theta)^{-\frac{1}{2}} \int_0^\infty \frac{1}{\sqrt{2\pi u}} e^{-\frac{1}{2}u} du$	B1	
		$=(1-2\theta)^{-\frac{1}{2}}$		Answer given
			[6]	
	(ii)	Either expand the mgf as a power series:		
		$(1-2\theta)^{-\frac{1}{2}} = 1+\theta+3\frac{\theta^2}{2!}+\dots$	M1A1	
		$E(Y) = 1$ (coefficient of $\theta$ )	B1	
		$E(Y^2) = 3$ (coefficient of $\theta^2/2!$ )	B1 B1	
		Hence $Var(Y) = 2$		
			[5]	
		Or differentiate the mgf:	M1	
		$1^{\text{st}}$ derivative simplifies to $(1-2\theta)^{-\frac{3}{2}}$	A1	
		Putting $\theta = 0$ gives $E(Y) = 1$	B1	
		$2^{\text{nd}}$ derivative simplifies to $3(1-2\theta)^{-\frac{5}{2}}$	A1	
		Putting $\theta = 0$ gives $E(Y^2) = 3$ , hence $Var(Y) = 2$	B1	
			[5]	
	(iii)	For independent rvs <i>X</i> and <i>Y</i> ,		
		$M_{X+Y}(\theta) = M_X(\theta) M_Y(\theta)$	B1	

Question	Answer	Marks	Guidance
	Hence $M_U(\theta) = (1 - 2\theta)^{-\frac{\eta}{2}}$	B1	
	E(U) = n	B1	
	Var(U) = 2n	B1	
		[4]	
(iv)	$\mathbf{M}_{W}(\theta) = \mathbf{E}\left(\exp\left(\theta\left(\frac{U-n}{\sqrt{2n}}\right)\right)\right)$	B1	
	$= \exp\left(-\frac{\theta n}{\sqrt{2n}}\right) \mathbb{E}\left(\exp\left(\frac{\theta}{\sqrt{2n}}U\right)\right)$	B1	Or by use of general linear transformation result
	$= \exp\left(-\frac{\theta n}{\sqrt{2n}}\right) \mathbf{M}_{U}\left(\frac{\theta}{\sqrt{2n}}\right)$	B1	
	$=\exp\left(-\frac{\theta n}{\sqrt{2n}}\right)\left(1-\frac{2\theta}{\sqrt{2n}}\right)^{-\frac{n}{2}}$	B1	
	Hence $ln(M_W(\theta))$ as given		
	Expanding $ln(M_w(\theta))$ gives		
	$-\sqrt{\frac{n}{2}} \theta + \frac{n}{2} \left( \sqrt{\frac{2}{n}} \theta + \frac{1}{2} \left( \sqrt{\frac{2}{n}} \theta \right)^2 + \text{terms of order } n^{-\frac{3}{2}} \right)$	M1A1	Award a maximum of 3 marks from here on if no account is taken of terms
	Convincing simplification to $\frac{1}{2}\theta^2$ + terms of order $n^{-\frac{1}{2}}$	B1	beyond $\theta^2$ .
	Hence tends to $\frac{1}{2}\theta^2$ .		(answer given)
	The mgf therefore tends to $\exp(\frac{1}{2}\theta^2)$ .	B1	Uniqueness of mgfs may be
	That is, the distribution of <i>W</i> tends to the standard Normal.	B1	implied
		[9]	
		[24]	

(a)	(i)	Sample means:		
		standard 107.2766, GM 113.6192	B1	
		Sample variances: standard 138.2243, GM 121.9372	B1	
		H <sub>0</sub> : $\mu_1 = \mu_2$ (the means of the underlying distributions) H <sub>0</sub> : $\mu_1 \neq \mu_2$	B1	Must be clear that the hypotheses refer to underlying means
		Test statistic: $\frac{113.6192 - 107.2766}{\sqrt{\frac{1382243}{30} + \frac{12193721}{26}}} = 2.08(0103)$	M1A1	
		Compare with z distribution Critical value, 2 tails, 5%, 1.960	B1 B1	
		So 2.08(0103) is just in the critical region  Hence reject H <sub>0</sub> and conclude that there is a difference in mean weight between standard and GM tomatoes	B1B1	
		Assumption:		
		the tomatoes may be regarded, in some sense, as random samples of their respective varieties.	B1 [10]	
	(ii)	Additional assumptions:  underlying Normality of the tomatoes' weights common variance in the underlying distributions  Given the sample sizes it seems safe to use the Normal distribution	B1 B1 E1	
		The Normal test is better in that it makes fewer assumptions.	E1 [4]	
<b>(b)</b>	(i)	The tomatoes should be chosen at random.	B1	
		The panel should not know which tomatoes are GM / standard.	B1	
		Making fine judgements on the appearances of tomatoes is unlikely to be reliable.	B1	Accept other sensible comments
	<b>b</b> )		Test statistic:  \[ \frac{113.6192 - 107.2766}{\sqrt{1382243} + \frac{12193721}{26}} = 2.08(0103) \]  Compare with z distribution  Critical value, 2 tails, 5%, 1.960  So 2.08(0103) is just in the critical region  Hence reject H₀ and conclude that there is a difference in mean weight between standard and GM tomatoes.  Assumption:  the tomatoes may be regarded, in some sense, as random samples of their respective varieties.  (ii) Additional assumptions:  underlying Normality of the tomatoes' weights common variance in the underlying distributions  Given the sample sizes it seems safe to use the Normal distribution.  The Normal test is better in that it makes fewer assumptions.  (i) The tomatoes should be chosen at random.  The panel should not know which tomatoes are GM / standard.  Making fine judgements on the appearances of tomatoes is unlikely to	Test statistic: $\frac{113.6192 - 107.2766}{\sqrt{\frac{1382243}{30} + \frac{12193721}{26}}} = 2.08(0103)$ Compare with z distribution Critical value, 2 tails, 5%, 1.960 So 2.08(0103) is just in the critical region Hence reject H <sub>0</sub> and conclude that there is a difference in mean weight between standard and GM tomatoes. Assumption: the tomatoes may be regarded, in some sense, as random samples of their respective varieties.  [10]  (ii) Additional assumptions: underlying Normality of the tomatoes' weights common variance in the underlying distributions Given the sample sizes it seems safe to use the Normal distribution. The Normal test is better in that it makes fewer assumptions.  [1]  (i) The tomatoes should be chosen at random. The panel should not know which tomatoes are GM / standard. Making fine judgements on the appearances of tomatoes is unlikely to

Question	Answer		Guidance		
(ii)	Rank sums:			Or Mann-Whitney	
	Standard 134, GM 76	M1A1		79, 21	
	H <sub>0</sub> : GM and std tomatoes have, on the whole, the same appearance		Hope for, but don't expect, a		
	H <sub>1</sub> : GM tomatoes have, on the whole, a better appearance than std tomatoes	B1	in location parameter for		
	(lower rank sum for GM indicates that the evidence is in the correct tail)	B1	underlying distributions of appearances		
	Wilcoxon rank sum test	B1	May be implied		
	Critical value for $m = n = 10, 1$ tail, 1% level is 74	B1		Critical value 19	
	The observed value of 76 is not in the critical region			Citical value 19	
	so accept H <sub>0</sub> . That is, conclude that there is insufficient evidence to suppose that, on the whole, GM tomatoes have a better appearance	B1			
	than standard tomatoes.	[7]			
		[24]			

Ç	uesti	on	Answer	Marks	Guida	nce
4			$Y_{ij}$ denotes the $j$ th observation (or measurement) in the $i$ th group $\mu$ is the underlying mean for whole population $\alpha_i$ denotes the population mean difference for the $i$ th group $\varepsilon_{ij}$ denotes the random error in the $ij$ th observation (or measurement)	B1 B1 B1 [4]	Accept just $Y_{ij}$ denotes the observations (measurements)  Withhold these marks if there is no reference to populations	In marking section (a), evidence of understanding is more important than precise wording.
		(ii)	$\varepsilon_{ij} \sim \text{ind N}(0, \sigma^2),$ where $\sigma^2$ is constant across groups.  The variance ratio test (F test) requires underlying Normality (and common variance).	B1B1 B1 B1 [4]	B1 ind N, B1 zero mean May be implied by later comment Common variance may be implied by earlier statement	
		(iii)	$H_0$ : all the $\alpha_i$ are zero $H_1$ : not all the $\alpha_i$ are zero If $H_0$ is accepted then we proceed on the basis that the group means are all equal. I.e. there are no 'treatment effects'. If $H_0$ is rejected the we proceed on the basis that the group means are not all equal. I.e. there are some 'treatment effects'.	B1 B1 E1 E1 [4]	Accept clear wording in terms of all group effects being zero, or all groups having same underlying distribution	$H_1$ : all the $\alpha_i$ are non-zero scores zero
	(b)	(i)	k = 4 (number of routes is 4) N = 19 (number of journeys is 19)	B1 B1 [2]		

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Question				Ansv	wer				Marks	S Guidance	
(i	ii)	Sc	ource of var'n	SS	df	MS	F		M1	Evidence of understanding how to complete the table  for the figures shown in bold -1 each error	
		Ве	etween groups	333.77	3	111.2567	3.98(113)		A4		
		W	ithin groups	419.19	15	27.9460					
		To	otal	752.96	18				A4		
		H <sub>1</sub> : rout The obs The 5% Observe That is,	outes have the sames do not all have the same erved variance rational value is 3. The ed value is in critical proceed on the assing mean time.	the same used io is comp 29 (tables al region,	under pared ) or 3 so rej	lying mean t with $F_{3, 15}$ .287382 (cal ject $H_0$ .	culator)		B1 M1 A1 B1 B1 [10] [24]	Hypotheses may be implied by later comments  May be implied by use of a correct value  Or equivalent wording	

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