



Mathematics (MEI)

Advanced GCE 4768

Statistics 3

Mark Scheme for June 2010

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r		1		
Q1	$D \sim N(2018, \sigma = 96)$		When a candidate's answers suggest that (s)he appears to have neglected to use the difference columns of the Normal distribution tables penalise the first occurrence only.	
(i)	Systematic Sampling. It lacks any element of randomness. Choose a random starting point in the range 1 – 10.	B1 E1 E1	May be implied by the next mark. Allow reasonable alternatives e.g. "the list may contain cycles." Beware proposals for a different sampling	[3]
			memod.	
(ii)	$P(D > 2100) = P\left(Z > \frac{2100 - 2018}{96} = 0.8542\right)$	M1 A1	For standardising. Award once, here or elsewhere.	
	= 1 - 0.8034 = 0.1966	A1	c.a.o.	[3]
(iii)	$D_1 + D_2 + D_3 \sim N(6054,$	B1	Mean.	
	$\sigma^2 = 96^2 + 96^2 + 96^2 = 27648)$	B1	Variance. Accept sd (= 166.277).	
	P(this < 6000) = P $\left(Z < \frac{6000 - 6054}{166.277} = -0.3248\right)$ = 1 - 0.6273 = 0.3727	Al	c.a.o.	
	Must assume that the months are independent.	E1	Reference to independence of months.	
	This is unlikely to be realistic since e.g. consecutive months may not be independent.	E1	Any sensible comment.	[5]
(iv)	Claim ~ N(2018 × 0.45 + 21200 × 0.10 = 3028.10,	M1	Mean.	
	$96^2 \times 0.45^2 + 1100^2 \times 0.10^2 = 13966.24$	A1 M1 A1	c.a.o. Variance. Accept sd (= 118.18). c.a.o.	
	$P(3000 < \text{this} < 3300) = P\left(\frac{3000 - 3028.1}{118.18} < Z < \frac{3300 - 3028.1}{118.18}\right)$	M1	Formulation of requirement: a two-sided inequality.	
	= P(-0.2378 < Z < 2.3008) = 0.9893 - (1 - 0.5940) = 0.5833	A1 A1	Ft c's parameters. c.a.o.	[7]
			Total	[19]
L		L	Total	[10]

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Q2						
(i)	 A <i>t</i> test might be used because sample is small population variance is unknown Must assume background population is Normal. 	B1 B1 B1		[3]		
(ii)	H ₀ : $\mu = 1.040$ H ₁ : $\mu \neq 1.040$ where μ is the mean specific gravity of the mixture.	B1 B1	Both hypotheses. Hypotheses in words only must include "population". Do NOT allow " $\overline{X} =$ " or similar unless \overline{X} is clearly and explicitly stated to be a <u>population</u> mean. For adequate verbal definition. Allow absence of "population" if correct notation μ is used.			
	$\overline{x} = 1.0452$ $s_{n-1} = 0.007155$ Test statistic is $\frac{1.0452 - 1.040}{0.007155}$	B1 M1	$s_n = 0.006746$ but do <u>NOT</u> allow this here or in construction of test statistic, but FT from there. Allow c's \overline{x} and/or s_{n-1} . Allow alternative: 1.040 + (c's 1.860) × $\frac{0.007155}{\sqrt{9}}$ (= 1.0444) for subsequent			
	= 2.189(60).	A1	comparison with \overline{x} . (Or $\overline{x} - (c's 860) \times \frac{0.007155}{\sqrt{9}}$ (= 1.0407) for comparison with 1.040.) c.a.o. but ft from here in any case if wrong. Use of 1.040 - \overline{x} scores M1A0, but ft.			
	 Refer to t₈. Double-tailed 10% point is 1.860. Significant. Seems mean specific gravity in the mixture does not meet the requirement. 	M1 A1 A1 A1	No ft from here if wrong. P(t > 2.1896) = 0.05996. No ft from here if wrong. ft only c's test statistic. ft only c's test statistic.	[9]		
(iii)	CI is given by $1.0452 \pm 2.306 \times \frac{0.007155}{\sqrt{9}}$	M1 B1 M1				
	= $1.0452 \pm 0.0055 = (1.039(7), 1.050(7))$ In repeated sampling, 95% of confidence intervals constructed in this way will contain the true population mean.	A1 E2	c.a.o. Must be expressed as an interval. ZERO/4 if not same distribution as test. Same wrong distribution scores maximum M1B0M1A0. Recovery to t_8 is OK. E2, 1, 0.			
			Total	[18]		

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Q3											
(a) (i)	Use paired data in order to eliminate differences between authorities.										[1]
(ii)	$H_0: m = 0$ $H_1: m > 0$ where <i>m</i> is the population median difference.						Both. Acc Adequate	ept hypoth definition	eses in wor of <i>m</i> to incl	rds. lude	
						I	populatio	11 .			
	D:00(A)			1	5	4	2 11	0	2 0		
	Dill (Al	ler - Belo	(e) 6	- <u> </u>	5 -	-4 -	5 11	8	2 9		
	Ran	k of diff	6	l	5	4 3	9	1	2 8		
						M1 For differences. ZERO in this section if differences not used.					
						M1	For ranks.				
						A1	FT from h	ere if rank	s wrong		
	$W_{-} = 1 + 3$	+4=8 (e)	or = 2 + 5 + 6	+7+8+9 =	37)	B1					
	Refer to ta	bles of Wi	lcoxon pair	ed (/single	e sample)	M1	No ft from	here if wr	ong.		
	statistic for	r n = 9.									
	Lower 5%	point is 8	(or upper is	$37 \text{ if } W_+$	used).	A1	i.e. a 1-tail	test. No f	t from here	if wrong.	
	Result is s	ignificant.				A1	ft only c's	test statist	ic.		
	Evidence suggests the percentage has been raised (or						ft only c's	test statist	IC.		[10]
	the whole)	•									
(b)	H_0 : Stock	market prio	ces can be i	nodelled t	by Benford	's Law.					
	H_1 : Stock	market prio	ces can not	be modell	led by Bent	ord's Lav	W.				
	Proh	0.201	0.176	0.125	0.007	0.070	0.067	0.058	0.051	0.046	
	Evn f	60.2	25.2	25.0	10.097	15.8	13.4	0.038	10.2	0.040	
	Exp I Obs f	55	24	23.0	19.4	15.0	15.4	11.0	10.2	9.2	
	OUS I	33	54	27	10	15	1 /	12	15	9	
	$X^{2} = 0.44917 + 0.04091 + 0.16 + 0.59588 + 0.04051 + 0.96716 + 0.01379 + 2.25882 + 0.00435 = 4.5305(9)$ Refer to χ_{8}^{2} .						M1Probs \times 200 for expected frequencies. All correct.M1Calculation of X^2 .				
							c.a.o.				
							M1 Allow correct df (= cells -1) from				
							no ft if wr	ouped tab	ie and it. C	ulerwise,	
							$P(X^2 > 4.5)$	3059) = 0.	80636.		
	Upper 5% point is 13.36.					Al	No ft from	here if wr	ong.		
	Not significant.					Al	tt only c's	test statist	1C.		
	Suggests Benford's Law provides a reasonable					AI	It only c's	test statist	1C.		[7]
	model in the context of share prices.										
						-				T ()	[10]
										Total	[18]

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Q4	$f(x) = \lambda e^{-\lambda x}$ for $x \ge 0$, where $\lambda > 0$.		Given $\int_0^\infty x^r e^{-\lambda x} dx = \frac{r!}{\lambda^{r+1}}$	
(i)	$\int_0^\infty f(x) dx = \int_0^\infty \lambda e^{-\lambda x} dx$ $= \left[-e^{-\lambda x} \right]_0^\infty$ $= \left(0 - (-e^0) \right) = 1$	M1 M1 A1	Integration of $f(x)$. Use of limits or the given result. Convincingly obtained (Answer given.)	
		G1 G1	Curve, with negative gradient, in the first quadrant only. Must intersect the <i>y</i> -axis. (0, λ) labelled; asymptotic to x-axis.	[5]
(ii)	$E(X) = \int_0^\infty \lambda x e^{-\lambda x} dx$	M1	Correct integral.	
	$= \lambda \frac{1}{\lambda^2} = \frac{1}{\lambda}$ $E(X^2) = \int_0^\infty \lambda x^2 e^{-\lambda x} dx$	A1 M1	c.a.o. (using given result) Correct integral.	
	$=\lambda \frac{2}{\lambda^3} = \frac{2}{\lambda^2}$	A1 M1	c.a.o. (using given result) Use of $F(X^2) = F(X)^2$	
	$\operatorname{Var}(X) = \operatorname{E}(X^{2}) - \operatorname{E}(X)^{2} = \frac{2}{\lambda^{2}} - \left(\frac{1}{\lambda}\right)^{2} = \frac{1}{\lambda^{2}}$	A1	Use of $E(X) = E(X)$	[6]
(iii)	$\mu = 6 \qquad \therefore \lambda = \frac{1}{6}$ $\overline{X} \sim (\text{approx}) \operatorname{N}\left(6, \frac{6^2}{50}\right)$	B1 B1 B1 B1	Obtained λ from the mean. Normal. Mean. ft c's λ . Variance. ft c's λ .	[4]
(iv)	<u>EITHER</u> can argue that 7.8 is more than 2 SDs from μ . $(6+2\sqrt{0.72} = 7.697;$ <u>must</u> refer to SD (\overline{X}), not SD(X))	M1	A 95% C.I would be (6.1369, 9.4631).	
	i.e. outlier. \Rightarrow doubt.	M1 A1		[3]
	OR formal significance test: $\frac{7.8-6}{\sqrt{0.72}} = 2.121$, refer to N(0,1), sig at (eg) 5% \Rightarrow doubt.	M1 M1 A1	Depends on first M, but could imply it. P($ Z > 2.121$)= 0.0339	
			Total	[18]

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